

Estimation

Introduction to Quantitative Social Science

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What is Statistical Inference?

- Guessing what we do not observe from what we do observe
- What we want to estimate: **parameter** $\theta \rightsquigarrow$ unobservable
- What you do observe: **data**
- We use data to compute an estimate of the parameter $\hat{\theta}$

- How good is $\hat{\theta}$ as an estimate of θ ?
- Ideally, we want to know **estimation error** $= \hat{\theta} - \theta_0$ where θ_0 is the true value of θ
- The problem: θ_0 is unknown
- Instead, we consider two hypothetical scenarios:
 - 1 How well would $\hat{\theta}$ perform as the sample size goes to infinity?
 - 2 How well would $\hat{\theta}$ perform over *repeated data generating process*?

Polling Disaster: 2016 Election

- All major preelection polls predicted Clinton's victory
- What happened?
 - FBI announcements
 - non-response bias
 - social desirability bias
 - failure to predict turnout
- We will look at polls closely in the precept this week
- For today, let's look at ABC News/Washington Post poll:
 - Nov. 3 – Nov. 6 (election was Nov. 8)
 - 2220 likely voters
 - Live phone
 - Clinton (47%), Trump (43%), Johnson (4%)
 - Margin of error: ± 2.5 percentage points
- Actual election result (national vote): Clinton (48%), Trump (47%)

Estimating Trump's Support

- Parameter θ : **population proportion** of likely voters who support Trump
 - Estimator $\hat{\theta}$: **sample proportion** of respondents who support Trump
 - How good is $\hat{\theta}$ as an estimate of θ ?
-
- Assume a simple random sampling of n voters: $n = 2220$
 - Define a random variable $X_i = 1$ if the i th respondent supports Trump and $X_i = 0$ otherwise for each $i = 1, 2, \dots, n$
 - Data generating process: Binomial distribution with success probability p and size n where p is the population proportion of likely voters who support Trump
 - Estimator: $\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$
 - That is, $\theta = p$ and $\hat{\theta} = \bar{X}$

- 1 How well would \bar{X} behave as the sample size increases?
 - Law of large numbers: $\bar{X} \rightarrow p$
 - **consistency**
 - But, how large is large enough?
- 2 How would \bar{X} behave over repeated data generating process?
 - hypothetical scenario: repeatedly conduct a survey under the exact same conditions many times
 - expectation = average performance: $\mathbb{E}(\bar{X}) = p$
 - **unbiasedness**
 - sampling distribution of \bar{X} : Binomial random variable divided by n
 - standard deviation of sampling distribution:

$$\sqrt{\mathbb{V}(\bar{X})} = \sqrt{\frac{p(1-p)}{n}}$$

- **standard error** = estimated standard deviation

$$\sqrt{\widehat{\mathbb{V}(\bar{X})}} = \sqrt{\frac{\bar{X}(1-\bar{X})}{n}} = \sqrt{\frac{0.43 \times (1-0.43)}{2200}} \approx 0.011$$

Confidence Intervals

- Beyond standard error: characterizing the whole sampling distribution
- Central limit theorem: for a sufficiently large sample size,

$$\bar{X} \stackrel{\text{approx.}}{\sim} \mathcal{N}\left(\mathbb{E}(X), \frac{\mathbb{V}(X)}{n}\right)$$

- In the current case:

$$\bar{X} \stackrel{\text{approx.}}{\sim} \mathcal{N}\left(p, \frac{p(1-p)}{n}\right)$$

- Choose the level of **confidence interval**: e.g., 95%
- Compute the confidence interval, which contains the true value, e.g., 95% of time over repeated data generating process

- $(1 - \alpha) \times 100\%$ (asymptotic) confidence intervals:

$$CI_{\alpha} = [\bar{X} - z_{\alpha/2} \times \text{standard error}, \bar{X} + z_{\alpha/2} \times \text{standard error}]$$

where $z_{\alpha/2}$ is called the **critical value**

- $P(Z > z_{\alpha/2}) = \alpha/2$ and $Z \sim \mathcal{N}(0, 1)$

- 1 $\alpha = 0.01$ gives $z_{\alpha/2} = 2.58$

- 2 $\alpha = 0.05$ gives $z_{\alpha/2} = 1.96$

- 3 $\alpha = 0.10$ gives $z_{\alpha/2} = 1.64$

- Be careful about the **interpretation!!**

- 1 Probability that the true value is in a *particular* confidence interval is either 0 or 1

- 2 Confidence intervals are *random*, while the truth is *fixed*

- CIs for ABC/WP poll:

$$90\%CI : [0.43 - 1.64 \times 0.011, 0.43 + 1.64 \times 0.011] = [0.412, 0.447]$$

$$95\%CI : [0.43 - 1.96 \times 0.011, 0.43 + 1.96 \times 0.011] = [0.409, 0.451]$$

$$99\%CI : [0.43 - 2.58 \times 0.011, 0.43 + 2.58 \times 0.011] = [0.402, 0.457]$$

Summary: Inference with Random Sampling

- Random sampling from a large population
- **Sample analogue** principle: use sample mean to infer population mean
- Asymptotic inference:

- 1 Law of large Numbers:

$$\bar{X} \rightsquigarrow \mathbb{E}(X)$$

- 2 Central Limit Theorem:

$$\bar{X} \stackrel{\text{approx.}}{\sim} \mathcal{N}\left(\mathbb{E}(X), \frac{\mathbb{V}(X)}{n}\right)$$

- Standard error: $\sqrt{\frac{\mathbb{V}(X)}{n}}$
- Confidence interval:

$$[\bar{X} - z_{\alpha/2} \times \text{standard error}, \bar{X} + z_{\alpha/2} \times \text{standard error}]$$

Comparison of Two Samples

- Comparison of two groups is more interesting
- Public opinion differences across groups
- Difference between treatment and control groups in experiments

- Causal inference with randomized experiments
- Back to the GOTV example
- The 2006 Michigan August primary experiment
- Treatment Group: postcards showing their own and their neighbors' voting records
- Control Group: received nothing

Social Pressure Experiment Revisited

- Turnout rate: $\bar{X}_T = 0.37$, $\bar{X}_C = 0.30$,
- Sample size: $n_T = 360$, $n_C = 1890$
- Estimated **average treatment effect**:

$$\widehat{ATE} = \bar{X}_T - \bar{X}_C = 0.07$$

- Standard error:

$$\sqrt{\frac{\bar{X}_T(1 - \bar{X}_T)}{n_T} + \frac{\bar{X}_C(1 - \bar{X}_C)}{n_C}} = 0.028$$

- 95% Confidence intervals based on CLT:

$$\begin{aligned} & [\widehat{ATE} - \text{standard error} \times z_{0.025}, \widehat{ATE} + \text{standard error} \times z_{0.025}] \\ & = [0.016, 0.124] \end{aligned}$$

Minimum Wage Study Revisited

- Three identification strategies
 - 1 Cross-section comparison
 - 2 Before-and-after comparison
 - 3 Difference-in-differences
- How should we calculate the standard error under each strategy?
- What about confidence intervals?

```
minwage <- read.csv("data/minwage.csv")
## proportion of those fully employed before and after
## the increase in the minimum wage
minwage$fullPropBefore <- minwage$fullBefore /
  (minwage$fullBefore + minwage$partBefore)
minwage$fullPropAfter <- minwage$fullAfter /
  (minwage$fullAfter + minwage$partAfter)
## separate NJ and PA
minwageNJ <- subset(minwage, subset = (location != "PA"))
minwagePA <- subset(minwage, subset = (location == "PA"))
```

Cross-section Comparison: Assume no confounder

- Estimate: $\widehat{ATE} = \bar{X}_{NJ} - \bar{X}_{PA}$

```
est <- mean(minwageNJ$fullPropAfter) -  
      mean(minwagePA$fullPropAfter)  
est  
## [1] 0.0481
```

- Standard error:

$$\sqrt{\frac{\widehat{V}(X_{NJ})}{n_{NJ}} + \frac{\widehat{V}(X_{PA})}{n_{PA}}}$$

```
nNJ <- nrow(minwageNJ)  
nPA <- nrow(minwagePA)  
se <- sqrt(var(minwageNJ$fullPropAfter) / nNJ +  
           var(minwagePA$fullPropAfter) / nPA)  
se  
## [1] 0.0336
```

- Confidence intervals based on CLT:

$$[\widehat{\text{ATE}} - \text{standard error} \times z_{\alpha/2}, \widehat{\text{ATE}} + \text{standard error} \times z_{\alpha/2}]$$

```
## 90%
c(est - se * qnorm(0.95), est + se * qnorm(0.95))
## [1] -0.00715  0.10338

## 95%
c(est - se * qnorm(0.975), est + se * qnorm(0.975))
## [1] -0.0177  0.1140

## 99%
c(est - se * qnorm(0.995), est + se * qnorm(0.995))
## [1] -0.0384  0.1347
```

- Conservative inference based on Student's t -distribution is possible
- Comparison of two sample means from Normal distributions
- No exact distribution exists \rightsquigarrow approximation

```
t.test(minwageNJ$fullPropAfter, minwagePA$fullPropAfter)

##
## Welch Two Sample t-test
##
## data:  minwageNJ$fullPropAfter and minwagePA$fullPropAfter
## t = 1, df = 100, p-value = 0.2
## alternative hypothesis: true difference in means is not equal to 0
## 95 percent confidence interval:
##  -0.0185  0.1148
## sample estimates:
## mean of x mean of y
##      0.320      0.272
```

Before-and-After Comparison

- Assumption: only change is the treatment
- Estimate: $\widehat{ATE} = \bar{X}_{NJ,after} - \bar{X}_{NJ,before}$

```
est <- mean(minwageNJ$fullPropAfter) -  
      mean(minwageNJ$fullPropBefore)  
est  
## [1] 0.0239
```

- Standard error: $\sqrt{\widehat{V(ATE)}}$
- Variance of the sum of random variables:

$$\begin{aligned}V(X + Y) &= V(X) + V(Y) + 2Cov(X, Y) \\V(aX + bY) &= a^2V(X) + b^2V(Y) + 2abCov(X, Y)\end{aligned}$$

- Variance of \widehat{ATE} :

$$\begin{aligned}\mathbb{V}(\widehat{ATE}) &= \mathbb{V}(\bar{X}_{NJ,after}) + \mathbb{V}(\bar{X}_{NJ,before}) - 2\text{Cov}(\bar{X}_{NJ,after}, \bar{X}_{NJ,before}) \\ &= \frac{\mathbb{V}(X_{NJ,after})}{n_{NJ}} + \frac{\mathbb{V}(X_{NJ,before})}{n_{NJ}} - \frac{2\text{Cov}(X_{NJ,after}, X_{NJ,before})}{n_{NJ}}\end{aligned}$$

- Standard error:

```
se <- sqrt((var(minwageNJ$fullPropAfter) +  
            var(minwageNJ$fullPropBefore) -  
            2 * cov(minwageNJ$fullPropAfter,  
                    minwageNJ$fullPropBefore)) / nNJ)  
se  
## [1] 0.0176
```

- 95% confidence interval:

```
c(est - se * qnorm(0.975), est + se * qnorm(0.975))  
## [1] -0.0107 0.0585
```


Difference-in-Differences

- Assumption: parallel trend assumption
- Estimate:

$$\widehat{ATE} = (\bar{X}_{NJ,after} - \bar{X}_{NJ,before}) - (\bar{X}_{PA,after} - \bar{X}_{PA,before})$$

```
est <- (mean(minwageNJ$fullPropAfter) -  
        mean(minwageNJ$fullPropBefore)) -  
        (mean(minwagePA$fullPropAfter) -  
         mean(minwagePA$fullPropBefore))  
est  
## [1] 0.0616
```

- Variance:

$$\mathbb{V}(\widehat{ATE}) = \mathbb{V}(\bar{X}_{NJ,after} - \bar{X}_{NJ,before}) + \mathbb{V}(\bar{X}_{PA,after} - \bar{X}_{PA,before})$$

- Standard error:

```
se <- sqrt(var(minwageNJ$fullPropAfter) / nNJ +
           var(minwageNJ$fullPropBefore) / nNJ -
           2 * cov(minwageNJ$fullPropAfter,
                  minwageNJ$fullPropBefore) / nNJ +
           var(minwagePA$fullPropAfter) / nPA +
           var(minwagePA$fullPropBefore) / nPA -
           2 * cov(minwagePA$fullPropAfter,
                  minwagePA$fullPropBefore) / nPA)
se
## [1] 0.0455
```

- 95% confidence interval:

```
c(est - se * qnorm(0.975), est + se * qnorm(0.975))
## [1] -0.0276 0.1508
```